

# Inference for non-Stationary, non-Gaussian, Irregularly-Sampled Processes

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- 1 Introduction  
Easy Time Series Examples
- 2 Extension 1: Irregular Sampling
- 3 Extension 2: Non-Gaussian
- 4 Extension 3: Non-Stationary
- 5 Wrap-up

# Outline

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# Basic Ideas

The entire talk in one slide:

- Basic idea 1: Even if we only observe events  $Y_i$  at discrete times  $\{t_i\}$ , we can *model* these as observations  $Y(t_i)$  of a *process*  $Y(t)$  that is well-defined (if un-observed) in **continuous time**. This overcomes the “regularly-spaced observations” limitation.

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- Basic idea 2: Almost anything you can do with Gaussian distributions can also be done with other **Infinitely-Divisible (ID) distributions**, such as Poisson, Negative Binomial, Gamma,  $\alpha$ -Stable, Cauchy, Beta Process. Some of these feature discrete (integer) values, and some feature heavy tails, offering a wider range of behavior than is possible with Gaussian time series or processes.

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- Basic idea 3: **Start with Stationary** processes (like the Ornstein-Uhlenbeck process) or processes with Stationary increments (like Brownian Motion), **then extend** to *non-stationary* processes (like Diffusions).

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# Easy TS Example: AR(1)

Fix  $\mu \in \mathbb{R}$ ,  $\sigma^2 > 0$ , and  $\rho \in (0, 1)$ .

Draw  $X_0 \sim \text{No}(\mu, \sigma^2)$ .

For times  $t = 1, 2, \dots$ , set

$$\begin{aligned} X_t &:= \rho X_{t-1} + \zeta_t \\ &= \rho^t X_0 + \sum_{s=0}^{t-1} \zeta_{t-s} \rho^s = \sum_{s=0}^{\infty} \zeta_{t-s} \rho^s \end{aligned}$$

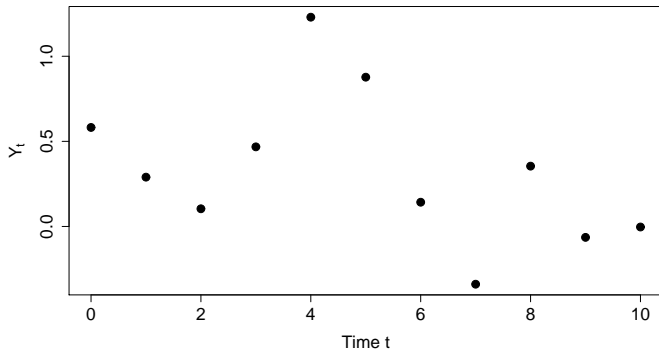
with Normal innovations  $\{\zeta_t\} \stackrel{\text{iid}}{\sim} \text{No}((1 - \rho)\mu, (1 - \rho^2)\sigma^2)$

Like most commonly-studied Time Series models, this features:

- **Regularly-spaced** observation times  $t = 0, 1, 2, 3, \dots$ ;
- **Gaussian** marginal distributions  $X_t \sim \text{No}(\mu, \sigma^2)$ ;
- **Stationary** distributions, with autocorrelation  $\text{Corr}(X_s, X_t) = \rho^{|t-s|}$ .



### Gaussian AR(1) Time Series: Equally-Spaced



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## Extension 1: Continuous Time

Set  $\lambda := -\log \rho > 0$  and let  $\zeta(ds)$  be a random measure assigning to disjoint intervals  $(a_i, b_i]$  independent random variables

$$\zeta((a, b]) \sim \text{No}((b-a)\lambda\mu, (b-a)2\lambda\sigma^2).$$

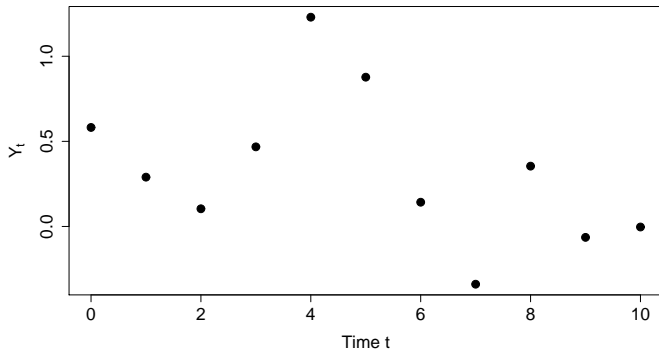
Now take  $X_0 \sim \text{No}(\mu, \sigma^2)$  and for  $t > 0$  set

$$X_t := X_0 e^{-\lambda t} + \int_0^t e^{-\lambda(t-s)} \zeta(ds) = \int_{-\infty}^t e^{-\lambda(t-s)} \zeta(ds)$$

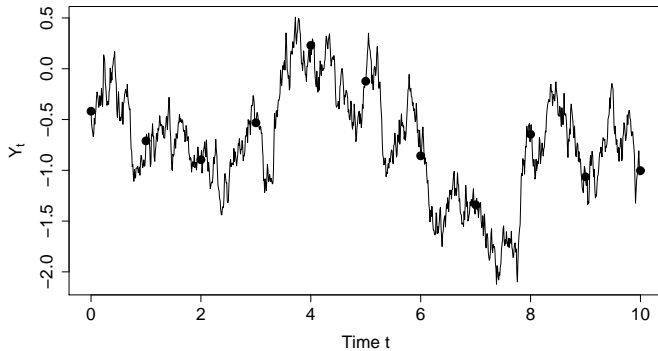
to find a process *for all*  $t > 0$  (or all  $t \in \mathbb{R}$ ) with the exact same joint distribution at  $t \in \mathbb{N}_0 = \{0, 1, 2, \dots\}$ , with

- **Irregularly-spaced** observation times  $\{t_i\} \subset \mathbb{R}_+$ ;
- **Gaussian** marginal distributions  $X_t \sim \text{No}(\mu, \sigma^2)$ ;
- **Stationary** distributions, with autocorrelation  $\text{Corr}(X_s, X_t) = \rho^{|t-s|}$ .

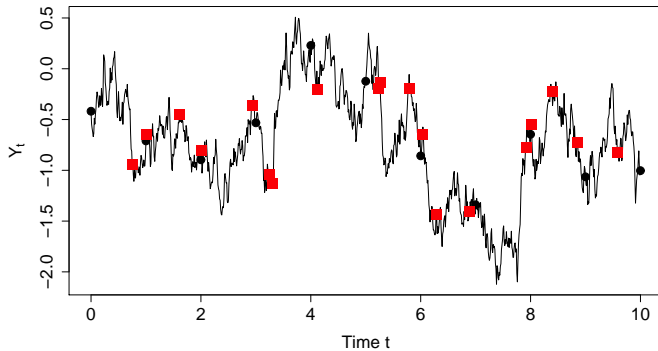
### Gaussian AR(1) Time Series: Equally-Spaced



### O-U Process with Regularly-Spaced Observations



### O-U Process with Irregularly-Spaced Observations



## Inference 1: Continuous Time

Inference is easy— MLEs or conjugate Bayes estimators are readily available, as is LH for arbitrary observation pairs  $\{(X_{t_i}, t_i)\}$  because  $s < t \Rightarrow$

$$X_t \mid \{X_u : u \leq s\} \sim \text{No}(\mu + (X_s - \mu)\rho, \sigma^2(1 - \rho^2)), \quad \rho := e^{-\lambda|t-s|}$$

so with Metropolis-Hastings we can sample posterior for any prior  $\pi(\mu, \sigma^2, \lambda)$ .

Easy extension to AR( $p$ ):

$$X_t + a_1 X_{t-1} + \dots + a_p X_{t-p} := \zeta_t$$

For example, by expressing as a **vector** AR(1) for

$$\mathbf{X}_t = [X_t, X_{t-1}, \dots, X_{t-p+1}]'$$

$$\mathbf{X}_t := R\mathbf{X}_{t-1} + \zeta_t$$

for some  $a \in \mathbb{R}^p$

## Extension 1: Continuous Time, More Broadly

Any stationary  $AR(p)$  has a  $MA(\infty)$  representation:

$$\rho(L)X(t) = \zeta_t$$

for the left-shift operator  $LX(t) := X(t - 1)$  and the polynomial

$$\rho(z) = 1 + \sum_{j=1}^p a_j z^j = 1 + a_1 z + \cdots + a_p z^p.$$

If  $\rho(z)$  has all its roots outside the unit circle, then

$$\frac{1}{\rho(z)} = 1 + \sum_{i=1}^{\infty} b_i z^i = 1 + b_1 z + b_2 z^2 + \cdots$$

$$X(t) = \frac{1}{\rho}(L)(\zeta_t) = \zeta_t + b_1 \zeta_{t-1} + b_2 \zeta_{t-2} + \cdots$$

Under **suitable conditions**, this has a continuous version

$$X(t) = \int_{-\infty}^t b(t-s)\zeta(ds)$$

for all  $t \geq 0$ , with the same distribution at times  $t \in \mathbb{N}_0$ .



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## Extension 2: Non-Gaussian

Just as **AR( $\rho$ ) Time Series**

$$X_t = \rho^t X_0 + \sum_{s=0}^{t-1} \zeta_{t-s} \rho^s$$

can be constructed with *any* iid innovations  $\{\zeta_s\}$ , so too for the continuous-time version

$$X_t := X_0 e^{-\lambda t} + \int_0^t e^{-\lambda(t-s)} \zeta(ds).$$

The **random measure**  $\zeta(ds)$  can be anything that is “iid” in the sense that:

- For disjoint sets  $A_i \subset \mathbb{R}$ , the RVs  $\{\zeta(A_i)\}$  are indep;
- The  $\{\zeta(a, b]\}$  distributions only depend on  $(b - a)$ .

For example, can have

$$\zeta(a, b] \sim \text{No}(\mu(b - a), \sigma^2(b - a))$$

$$\zeta(a, b] \sim \text{Ga}(\alpha(b - a), \beta)$$

$$\zeta(a, b] \sim \text{St}_A(\alpha, \beta, \gamma(b - a), \delta(b - a))$$

Ornstein-Uhlenbeck

Positive, Expo. Tails

Heavy (Pareto) Tails

## Extension 2: Non-Gaussian

The two conditions

- For disjoint sets  $A_i \subset \mathbb{R}$ , the RVs  $\{\zeta(A_i)\}$  are indep;
- The  $\{\zeta(a, b)\}$  distributions only depend on  $(b - a)$ .

require that each  $\zeta(A)$  should be **Infinitely Divisible**, or **ID**.

Examples:

ID Continuous	ID Discrete	Not ID
Normal	Poisson	Binomial
Gamma	Negative Binomial	Beta
$\alpha$ -Stable	$p_i \propto \frac{1}{(i+a)(i+b)}$	Uniform

## Extension 2: Non-Gaussian

But what about **Autocorrelated Count Data**?

- Binned photon counts in satellite Gamma Ray detectors?
- Binned photon counts in particle accelerator detectors?
- Seismic event counts?
- Pyroclastic flow counts?
- Rockfall counts near active volcano?
- Failures of complex systems?
- Rare disease case counts?

AR( $\rho$ ) and its continuous versions wouldn't respect **Integer Nature** of data. If  $X_t \in \mathbb{Z}$ , and  $|\rho| < 1$ , then

$$X_{t+1} = \rho X_t + \zeta_{t+1} \notin \mathbb{Z}.$$

Alternatives?

## Extension 2: Non-Gaussian

Here's a way to construct, and make inference in,

- **stationary** (for now) process  $X_t$  with
- **continuous time**  $t \in \mathbb{R}_+$  for *any*
- **non-Gaussian** marginal **Infinitely Divisible (ID)** dist'ns including both continuous dist'ns ( $\text{Ga}(\theta, \beta)$ ,  $\text{St}_A(\alpha, \beta, \theta, \delta)$ , *etc.*) and **discrete count distributions** ( $\text{Po}(\theta)$ ,  $\text{NB}(\theta, p)$ , *etc.*)

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In fact, except for the special cases of Gaussian and Poisson, we can do this in **many ways**.

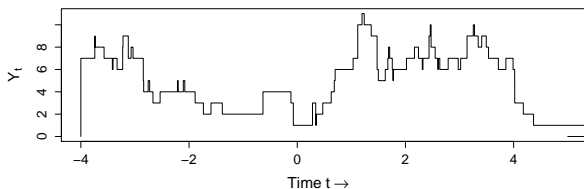
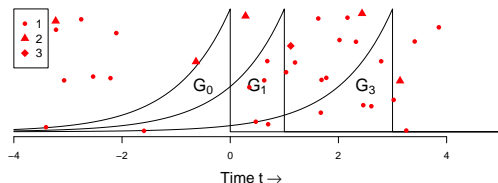
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In fact, except for the special cases of Gaussian and Poisson, we can do this in **many ways**. Let's look at an example.

## Ex 2: Negative Binomial Example



AR(1)-like Negative Binomial Process

Based on **Random Measure**  $\mathcal{N}(dx dy) \sim \text{NB}(\alpha dx dy, \beta)$  on  $\mathbb{R}^2$



# AR(1)-like Negative Binomial Process

Properties:

- $X_t \sim \text{NB}(\alpha, \beta)$  for all  $t$ ;
- $\text{Corr}(X_s, X_t) = \exp(-\lambda|t - s|)$ .

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Does this characterize the joint distribution of  $\{X_t\}$ ?

Nope. Here's a different process, the *Branching* NB:

- **Immigration** at rate  $\iota$ ;
- **Birth** at rate  $\beta$ ;
- **Death** at rate  $\delta$ ;

$$X_{t+\epsilon} = \begin{cases} X_t + 1 & \text{with probability } o(\epsilon) + \epsilon(\iota + \beta X_t) \\ X_t & \text{with probability } o(\epsilon) + 1 - \epsilon(\iota + (\beta + \delta)X_t) \\ X_t - 1 & \text{with probability } o(\epsilon) + \epsilon\delta X_t \end{cases}$$

$\sim \text{NB}(\alpha, \rho)$ , with autocovariance  $\exp(-\lambda|t - s|)$ .

# AR(1)-like Negative Binomial Process

Are these two processes the same?

# AR(1)-like Negative Binomial Process

Are these two processes the same?

Nope.

- The **Random Measure NB** process has jumps of all possible non-zero magnitudes  $\Delta X_t := [X_t - X_{t-}] \in \mathbb{Z} \setminus \{0\}$ , while the **Branching NB** process has only jumps of  $\Delta X_t = \pm 1$ ;
- The **Branching NB** process is Markov, so for  $s < t$

$$P[X_t \in A \mid \mathcal{F}_s] = P[X_t \in A \mid X_s],$$

while the **Random Measure NB** process isn't: if it has a jump  $\Delta X_s = 7$ , for example, then sooner or later there must follow a jump  $\Delta X_t = -7$ , so  $P[X_t = 3 \mid X_s = 10]$  depends on the history  $\mathcal{F}_s$ , not just the value  $X_s$ .

- The **Random Measure NB** process is **multivariate** ID, while **Branching NB** process is not.

# AR(1)-like Negative Binomial Process

Are these the only two?

# AR(1)-like Negative Binomial Process

Are these the only two?

Oh no.

- A discrete time AR(1)-like Markov process exists for **any ID distribution** and any “auto-correlation”  $\rho$ , based on **Thinning**:

$$Z \sim f(z | \lambda, \phi) = X + Y,$$

$$X \sim f(x | \rho\lambda, \phi) \perp\!\!\!\perp Y \sim f(y | \bar{\rho}\lambda, \phi) \quad [\bar{\rho} := (1 - \rho)]$$

$$X | Z \sim f(x | z, \rho, \lambda, \phi) = f(x | \rho\lambda, \phi) f(z - x | \bar{\rho}\lambda, \phi) / f(z | \lambda, \phi)$$

- Given  $\{X_m : m \leq n\} \sim f(x | \lambda, \phi)$ , draw

$$\xi_{n+1} \sim f(\xi | X_n, \rho, \lambda, \phi) \perp\!\!\!\perp \eta_{n+1} \sim f(\eta | \bar{\rho}\lambda, \phi)$$

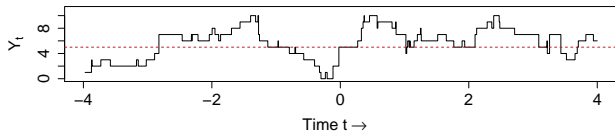
and set

$$X_{n+1} := \xi_{n+1} + \eta_{n+1} \sim f(x | \lambda, \phi).$$

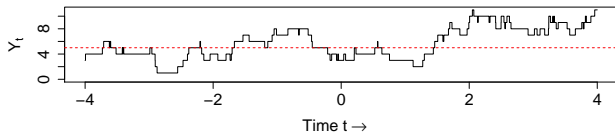
- A continuous-time version exists too.

# AR(1)-like Negative Binomial Process

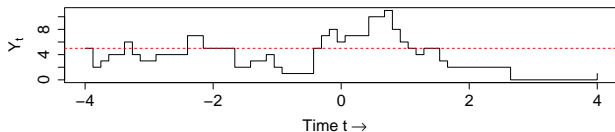
Random Measure NB( $\alpha = 10$ ,  $p = 0.67$ ,  $\lambda = 1$ )



Branching NB( $\alpha = 10$ ,  $p = 0.67$ ,  $\lambda = 1$ )



Continuous Thin NB( $\alpha = 10$ ,  $p = 0.67$ ,  $\lambda = 1$ )





# AR(1)-like Negative Binomial Process

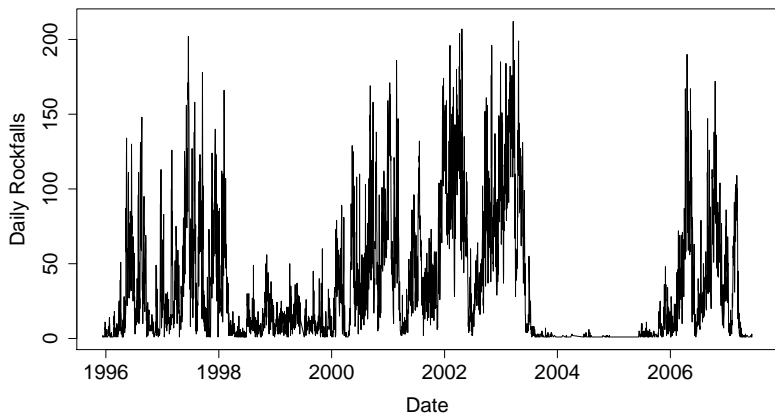
Larry Brown (U Penn) & I found six different AR(1)-like (also AR( $p$ )-like) processes for each ID distribution, with subtly different properties:

- Markov?
- Are *all* finite-dimensional marginals ID?
- In continuous time, are paths continuous? Increments  $\pm 1$ ? Or bigger?
- Time-reversible?

In Wolpert & L. D. Brown (2016+) we present a complete class theorem characterizing all *Markov Infinitely-Divisible Stationary Time-Reversible Integer-Valued* (“MISTI”) Processes.

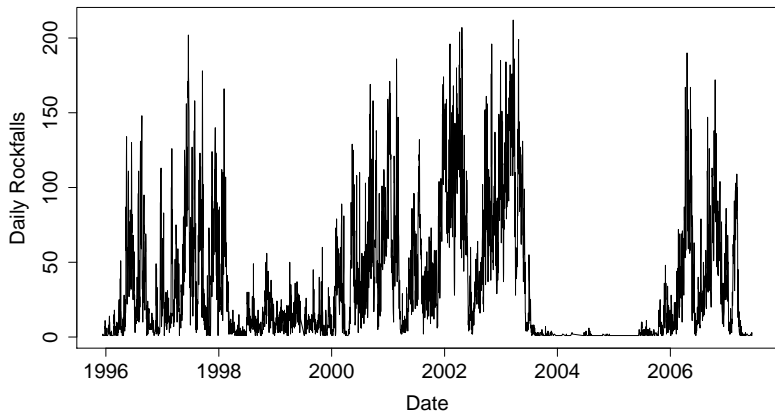
Let's look at a motivating example (not Astro, sorry!).

## Motivation for a Negative Binomial series...



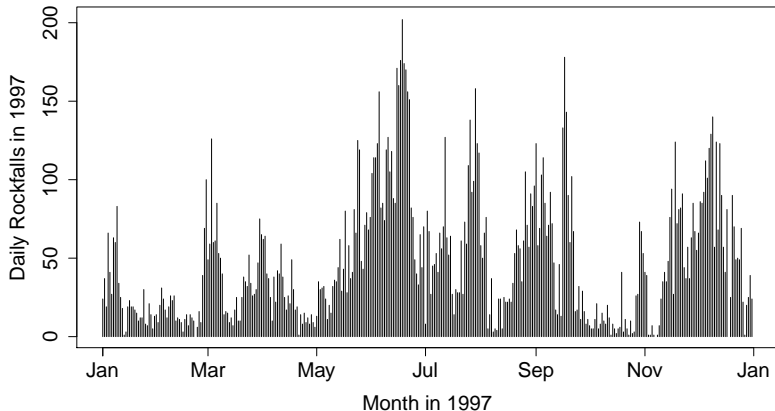
Rockfall counts at Soufrière Hills Volcano, Montserrat

## Stationary?



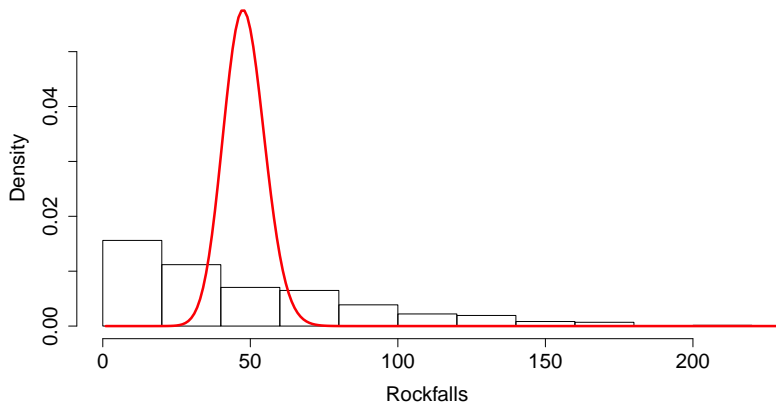
Looks a little patchy; maybe we can find a homogeneous subset. Looks like 1997 might be a less patchy year...

## Subset, just the 1997 Rockfalls



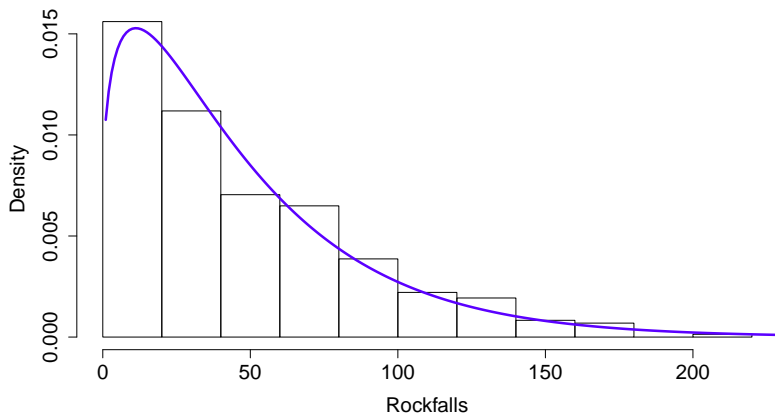
# Poisson fit is horrible...

Histogram of 1997 Rockfall Counts with  $Po(\lambda = 48.01)$  Overlay

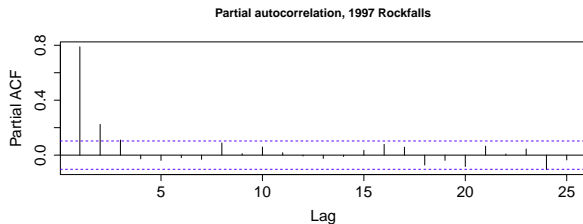
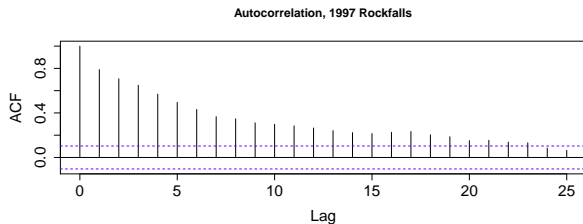


# Negative Binomial fit is great.

Histogram of 1997 Rockfall Counts with NB( $\alpha = 1.32$ ,  $p = 0.03$ ) Overlay



# But not IID, need autocorrelation



SO,

We (esp. **Jiangu Wang** '13) built a regression model based on one of the AR(1)-like NB models above, to relate

→ (easily counted) rockfall counts to

→ (hard to measure) subsurface volcanic magma flows,

to improve volcanic hazard forecasting.

Related issues were addressed by **Mary Beth Broadbent** ('14) in constructing NPB light curve models for Gamma Ray Bursts, with the help of astronomers **Tom Loredo** and **Jon Hakkila** (Thanks!).



## A shout out about **Another cool modeling idea:**

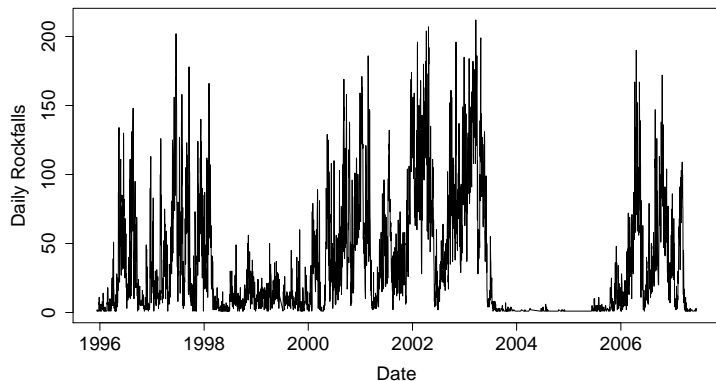
Robert Lund (Clemson) and his student Yunwei Cui (2009) found an interesting new way to model stationary integer-valued time series and processes using [Renewal Theory](#). In discrete time their method (**unlike ours**) is able to model [negative autocorrelation](#) and cyclic behavior. If you have negatively-autocorrelated integer data, look into it.

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## Extension 3: Non-Stationary

You might have noticed that the rockfall counts don't look stationary



So, we are implementing **random time-change**.

## Extension 3: Non-Stationary

Begin with stationary process  $X(t)$  ...

Then, construct **random time-change** model  $t \rightarrow R_t$  and set

$$Y_t := X(R_t)$$

We used:

$$R_t = \int_0^t \lambda(s) ds,$$

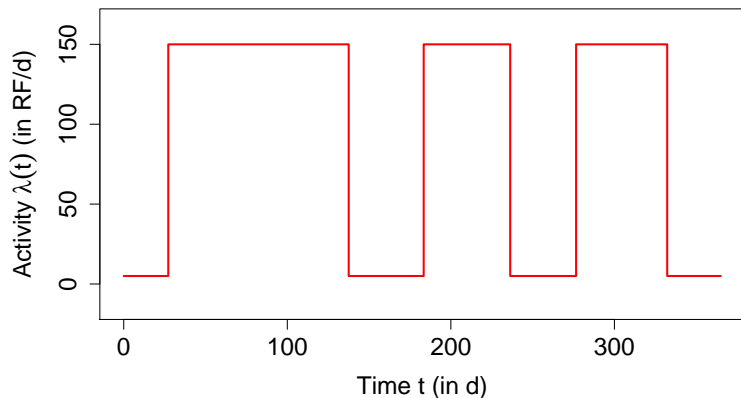
with two-level

$$\lambda(s) = \begin{cases} \lambda_+ & s_i < s \leq t_i \\ \lambda_- & t_i < s \leq s_{i+1} \end{cases}$$

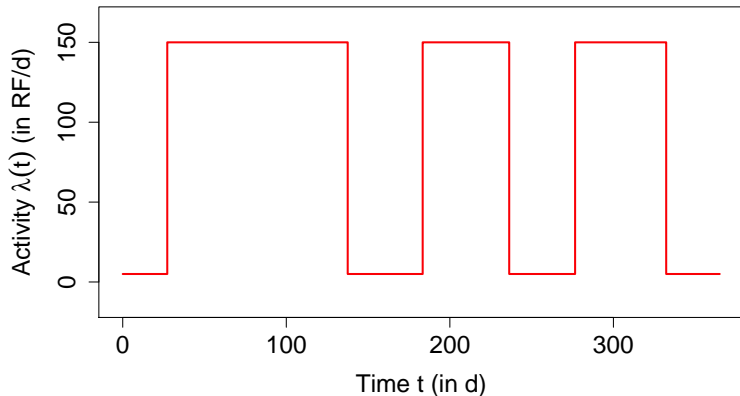
for uncertain levels  $0 < \lambda_- < \lambda_+ < \infty$  and transition times

$s_1 < t_1 < s_2 < t_2 < \dots$

## Extension 3: Non-Stationary



## Extension 3: Non-Stationary



A work in progress, wish us luck.

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# Wrap-up

## Conclusions:

- Statistical methods are available for data that are **non-Stationary**, or **non-Gaussian**, or **irregularly sampled**, or all three.
- They're not built into SAS.
- BUT, together, Astronomers and Statisticians can build problem-specific tools to support *estimation* and *prediction* and (especially Bayesian) *inference*,
- using routine simulation-based computational methods.



# Thanks!

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## Statistical, Mathematical and Computational Methods for Astronomy (ASTRO)

Glad to see you here!

